# Mathematical Statistics Notes

## Zoe Farmer

February 26, 2024

# 1 Preliminaries

## 1.1 Expected Value

The expected value for a discrete distribution is defined as

$$E[X] = \sum_{x} x \cdot P(X = x) = \sum_{x} x \cdot f(x)$$

For a continuous distribution, the expected value is defined as

$$E[X] = \int_{-\infty}^{\infty} x \cdot f(x)$$

The following properties hold.

1. Expectation is linear.

$$E[\alpha X + \beta Y] = \alpha E[x] + \beta E[y]$$

2. If X and Y are independent

$$E[XY] = E[X] \cdot E[Y]$$
$$E[g(X) \cdot h(Y)] = E[g(X)] \cdot E[h(Y)]$$

### 1.2 Variance

If we use  $\mu$  to denote the mean E[X], then the variance of X is defined by

$$Var[X] = E[(X - \mu)^{2}] = E[X^{2}] - (E[X])^{2}$$

The following properties hold.

- 1.  $\operatorname{Var}[aX] = a^2 \operatorname{Var}[x]$
- 2.  $Cov(X,Y) = E[(X \mu_X)(Y \mu_Y)] = E[XY] E[X]E[Y]$
- 3.  $\operatorname{Var}[X+Y] = \operatorname{Var}[X] + \operatorname{Var}[Y] + 2\operatorname{Cov}[X,Y]$

#### **1.3 Indicator Functions**

Instead of defining distributions piecewise as we've done in the past we prefer to use indicator functions that take on the values of zero and one.

Let A be a set. The function

$$I_A(x) = \begin{cases} 1 & \text{if } x \in A \\ 0 & \text{if } x \notin A \end{cases}$$

# 2 Four Important Tools

#### 2.1 Finding Distributions of Transformations of Random Variables

For discrete distributions it is enough to simply replace the variable with the function. This is most apparent through an example.

$$X \sim bin(n, p) \qquad Y = n - X$$
$$f_X(x) = \binom{n}{x} p^x (1-p)^{n-x}$$
$$f_Y(x) = \binom{n}{n-x} p^{n-x} (1-p)^y$$

For continuous distributions it's a little harder.

Let X be a continuous random variable with pdf  $f_x$ . Let Y be a random variable defined by Y = g(X) where g is invertible (and differentiable). Then the pdf for Y can be computed as

$$f_Y(y) = f_X\left(g^{-1}(y)\right) \cdot \left|\frac{d}{dy}g^{-1}(y)\right|$$

#### 2.2 Bivariate Transformations

Suppose that  $X_1$  and  $X_2$  are continuous random variables with joint pdf  $f_{X_1,X_2}(x_1,x_2)$  and suppose that new random variables  $Y_1$  and  $Y_2$  are defined by

$$Y_1 = g_1(X_1, X_2)$$
  $Y_2 = g_2(X_1, X_2)$ 

The joint pdf for  $Y_1$  and  $Y_2$  is given by

$$f_{Y_1,Y_2}(y_1,y_2) = f_{X_1,X_2}(g_1^{-1}(y_1,y_2),g_2^{-1}(y_1,y_2)) \cdot \left| \begin{array}{c} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_1}{\partial y_2} \end{array} \right|$$

If  $Y_1$  is a ratio, it is almost always a good idea to choose  $Y_2$  to be the denominator.

### 2.3 Order Statistics

We can define a short hand notation for the maximums and minimums.

The minimum is defined as

$$f_{X_{(1)}}(x) = n[1 - F(x)]^{n-1}f(x)$$

The maximum is defined as

$$f_{X_{(n)}}(x) = n[F(x)]^{n-1}f(x)$$

#### 2.4 Moment Generating Functions

For a random variable X, the moment generating function denoted by M(t) or  $M_X(t)$  is defined as

$$M(t) = E\left[e^{tX}\right]$$

We can use the Law of the Unconscious Statistician. Let X be a random variable with pdf  $f_x(x)$ . Let g(x) be some function. If X is discrete we have

$$E[G(X)] = \sum_{x} g(x) f_X(x)$$

If X is continuous we have

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) \, dx$$

In general, for a random variable X with mgf  $M_X(t)$ , the kth moment is obtained by  $M^{(k)}(0)$ , where  $M^{(k)}(t)$  is the kth derivative of  $M_X(t)$  with respect to t.

The moment generating function for a random variable X uniquely determines its distribution. If  $X_1, \ldots, X_n$  are iid random variables from a distribution with moment generating function  $M_X(t)$  then the sum  $Y = \sum_{i=1}^n X_i$  has moment generating function

$$M_Y(t) = \left[M_X(t)\right]^n$$

# 3 Estimators

# 4 Distributions

#### 4.1 The Gamma Distribution

#### 4.1.1 The Gamma Function

Defined for  $\alpha > 0$  as

$$\Gamma(\alpha) = \int_0^\infty x^{\alpha - 1} e^{-x} \, dx$$

Properties are as follows.

- 1.  $\Gamma(1) = 1$
- 2. For  $\alpha > 1$ ,  $\Gamma(\alpha) = (\alpha 1) \cdot \Gamma(\alpha 1)$
- 3. If  $n \ge 1$  is an integer,  $\Gamma(n) = (n-1)!$

#### 4.2 The Beta Distribution

#### 4.2.1 The Beta Function

The beta function is defined, for a, b > 0, as

$$\mathcal{B}(a,b) = \int_0^1 x^{a-1} (1-x)^{b-1} \, dx$$

The following property holds.

$$\mathcal{B}(a,b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$