# Mathematical Statistics Notes 

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## 1 Preliminaries

### 1.1 Expected Value

The expected value for a discrete distribution is defined as

$$
E[X]=\sum_{x} x \cdot P(X=x)=\sum_{x} x \cdot f(x)
$$

For a continuous distribution, the expected value is defined as

$$
E[X]=\int_{-\infty}^{\infty} x \cdot f(x)
$$

The following properties hold.

1. Expectation is linear.

$$
E[\alpha X+\beta Y]=\alpha E[x]+\beta E[y]
$$

2. If $X$ and $Y$ are independent

$$
\begin{aligned}
E[X Y] & = & E[X] \cdot E[Y] \\
E[g(X) \cdot h(Y)] & = & E[g(X)] \cdot E[h(Y)]
\end{aligned}
$$

### 1.2 Variance

If we use $\mu$ to denote the mean $E[X]$, then the variance of $X$ is defined by

$$
\operatorname{Var}[X]=E\left[(X-\mu)^{2}\right]=E\left[X^{2}\right]-(E[X])^{2}
$$

The following properties hold.

1. $\operatorname{Var}[a X]=a^{2} \operatorname{Var}[x]$
2. $\operatorname{Cov}(X, Y)=E\left[\left(X-\mu_{X}\right)\left(Y-\mu_{Y}\right)\right]=E[X Y]-E[X] E[Y]$
3. $\operatorname{Var}[X+Y]=\operatorname{Var}[X]+\operatorname{Var}[Y]+2 \operatorname{Cov}[X, Y]$

### 1.3 Indicator Functions

Instead of defining distributions piecewise as we've done in the past we prefer to use indicator functions that take on the values of zero and one.

Let $A$ be a set. The function

$$
I_{A}(x)= \begin{cases}1 & \text { if } x \in A \\ 0 & \text { if } x \notin A\end{cases}
$$

## 2 Four Important Tools

### 2.1 Finding Distributions of Transformations of Random Variables

For discrete distributions it is enough to simply replace the variable with the function. This is most apparent through an example.

$$
\begin{array}{r}
X \sim \operatorname{bin}(n, p) \quad Y=n-X \\
f_{X}(x)=\binom{n}{x} p^{x}(1-p)^{n-x} \\
f_{Y}(x)=\binom{n}{n-x} p^{n-x}(1-p)^{y}
\end{array}
$$

For continuous distributions it's a little harder.
Let $X$ be a continuous random variable with pdf $f_{x}$. Let $Y$ be a random variable defined by $Y=g(X)$ where $g$ is invertible (and differentiable). Then the pdf for $Y$ can be computed as

$$
f_{Y}(y)=f_{X}\left(g^{-1}(y)\right) \cdot\left|\frac{d}{d y} g^{-1}(y)\right|
$$

### 2.2 Bivariate Transformations

Suppose that $X_{1}$ and $X_{2}$ are continuous random variables with joint pdf $f_{X_{1}, X_{2}}\left(x_{1}, x_{2}\right)$ and suppose that new random variables $Y_{1}$ and $Y_{2}$ are defined by

$$
Y_{1}=g_{1}\left(X_{1}, X_{2}\right) \quad Y_{2}=g_{2}\left(X_{1}, X_{2}\right)
$$

The joint pdf for $Y_{1}$ and $Y_{2}$ is given by

$$
f_{Y_{1}, Y_{2}}\left(y_{1}, y_{2}\right)=f_{X_{1}, X_{2}}\left(g_{1}^{-1}\left(y_{1}, y_{2}\right), g_{2}^{-1}\left(y_{1}, y_{2}\right)\right) \cdot\left\|\begin{array}{cc}
\frac{\partial x_{1}}{\partial y_{1}} & \frac{\partial x_{1}}{\partial y_{2}} \\
\frac{\partial x_{2}}{\partial y_{1}} & \frac{\partial x_{1}}{\partial y_{2}}
\end{array}\right\|
$$

If $Y_{1}$ is a ratio, it is almost always a good idea to choose $Y_{2}$ to be the denominator.

### 2.3 Order Statistics

We can define a short hand notation for the maximums and minimums.

$$
\begin{array}{ccc}
X_{(1)} & = & \min \left(X_{1}, X_{2}, \ldots, X_{n}\right) \\
\vdots & & \vdots \\
X_{(n)} & = & \max \left(X_{1}, X_{2}, \ldots, X_{n}\right)
\end{array}
$$

The minimum is defined as

$$
f_{X_{(1)}}(x)=n[1-F(x)]^{n-1} f(x)
$$

The maximum is defined as

$$
f_{X_{(n)}}(x)=n[F(x)]^{n-1} f(x)
$$

### 2.4 Moment Generating Functions

For a random variable $X$, the moment generating function denoted by $M(t)$ or $M_{X}(t)$ is defined as

$$
M(t)=E\left[e^{t X}\right]
$$

We can use the Law of the Unconscious Statistician.
Let $X$ be a random variable with pdf $f_{x}(x)$. Let $g(x)$ be some function.
If $X$ is discrete we have

$$
E[G(X)]=\sum_{x} g(x) f_{X}(x)
$$

If $X$ is continuous we have

$$
E[g(X)]=\int_{-\infty}^{\infty} g(x) f_{X}(x) d x
$$

In general, for a random variable $X$ with $\operatorname{mgf} M_{X}(t)$, the $k$ th moment is obtained by $M^{(k)}(0)$, where $M^{(k)}(t)$ is the $k$ th derivative of $M_{X}(t)$ with respect to $t$.

The moment generating function for a random variable $X$ uniquely determines its distribution.
If $X_{1}, \ldots, X_{n}$ are iid random variables from a distribution with moment generating function $M_{X}(t)$ then the sum $Y=\sum_{i=1}^{n} X_{i}$ has moment generating function

$$
M_{Y}(t)=\left[M_{X}(t)\right]^{n}
$$

## 3 Estimators

## 4 Distributions

### 4.1 The Gamma Distribution

### 4.1.1 The Gamma Function

Defined for $\alpha>0$ as

$$
\Gamma(\alpha)=\int_{0}^{\infty} x^{\alpha-1} e^{-x} d x
$$

Properties are as follows.

1. $\Gamma(1)=1$
2. For $\alpha>1, \Gamma(\alpha)=(\alpha-1) \cdot \Gamma(\alpha-1)$
3. If $n \geq 1$ is an integer, $\Gamma(n)=(n-1)$ !

### 4.2 The Beta Distribution

### 4.2.1 The Beta Function

The beta function is defined, for $a, b>0$, as

$$
\mathcal{B}(a, b)=\int_{0}^{1} x^{a-1}(1-x)^{b-1} d x
$$

The following property holds.

$$
\mathcal{B}(a, b)=\frac{\Gamma(a) \Gamma(b)}{\Gamma(a+b)}
$$

